S-88.4200 Statistical Signal Processing. Final Exam May 23, 2011

- 1. Define or explain briefly the following concepts:
 - (a) Bias and variance of an estimator
 - (b) Consistency of estimates
 - (c) Efficiency of an estimator
 - (d) Fisher Information matrix
 - (e) Maximum Likelihood estimator
 - (f) Bayes Risk
- 2. Unbiased Estimator

Consider the i.i.d. data $\{x(0), x(2), \ldots, x(N-1)\}$, where each sample is distributed as $\mathcal{U}(0, a)$ (uniform distributed between 0 and a).

- (a) Find an unbiased estimator for a. The range of a is $0 < a < \infty$.
- (b) Show that the estimator derived under (a) is indeed an unbiased estimator for a.
- 3. Minimum Variance Unbiased Estimator
 - (a) Explain briefly the concept Minimum Variance Unbiased Estimator.
 - (b) Describe how Minimum Variance Unbiased Estimators are derived using concepts of sufficient statistics and completeness. Explain what is done in each step.
- 4. Suppose that $y_1, y_2, ..., y_N$ are independent samples of a Gaussian random variable with distribution $\mathcal{N}(\theta, \sigma_v^2)$ where θ is a random variable with a priori distribution $\mathcal{N}(a, \sigma_{\theta}^2)$. Assume that $\sigma_v^2, \sigma_{\theta}^2$, and a are known.
 - a) Find the Bayesian maximum a posteriori estimator $\hat{\theta}_{MAP}$ of θ .
 - b) Find the Bayesian estimator $\hat{\theta}_q$ of θ that minimizes the quadratic risk function.

5. A random parameter θ is uniformly distributed between (-10, 10). Find the Mean Square estimate of θ when we have made the 2 observations y(1) = -0.5 and y(2) = 1.0. The statistically independent measurements y(i) obey the distribution

$$f(y(i)|\theta) = \begin{cases} y(i) - \theta + 1, & \theta - 1 \le y(i) \le \theta \\ -y(i) + \theta + 1, & \theta < y(i) \le \theta + 1 \\ 0, & otherwise \end{cases}$$

Possibly useful formulas

The Bayes' rule:

$$f(\theta|\mathbf{y}) = \frac{f(\mathbf{y}|\theta)f(\theta)}{f(\mathbf{y})}$$

Binomial distribution

$$P_X(k) = \begin{cases} nkp^k(1-p)^{n-k} & k = 0, 1, 2, \dots, n \\ 0 & otherwise. \end{cases}$$

and $0 < \theta < 1$

Normal Distribution

$$f(y) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{1}{2\sigma^2}(y-\mu)^2}$$

Multivariate Gaussian with mean vector μ and covariance matrix **R**

$$f(\mathbf{y}) = \frac{1}{\sqrt{(2\pi)^N |R|}} exp \frac{-1}{2} [\mathbf{y} - \mu] \mathbf{R}^{-1} [\mathbf{y} - \mu]^{\mathrm{T}}$$

Conditional Mean when \mathbf{x} and \mathbf{y} are jointly Gaussian distributed

$$\mathrm{E}\left\{\mathbf{y}|\mathbf{x}\right\} = \mathrm{E}\left\{\mathbf{y}\right\} + \mathbf{R}_{\mathbf{y}\mathbf{x}}\mathbf{R}_{\mathbf{x}\mathbf{x}}^{-1}(\mathbf{x} - \mathrm{E}\left\{\mathbf{x}\right\}),$$

with

$$\mathbf{R}_{\mathbf{x}\mathbf{x}} = \mathbf{E}\left\{ (\mathbf{x} - \mathbf{E}\left\{\mathbf{x}\right\})(\mathbf{x} - \mathbf{E}\left\{\mathbf{x}\right\})^{\mathrm{T}} \right\},\label{eq:R_xx}$$

and

$$\mathbf{R}_{\mathbf{y}\mathbf{x}} = \mathrm{E}\left\{(\mathbf{y} - \mathrm{E}\left\{\mathbf{y}\right\})(\mathbf{x} - \mathrm{E}\left\{\mathbf{x}\right\})^{\mathrm{T}}\right\}.$$